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THE LAPLACE TRANSFORMATION OF THE DISTRIBUTION OF THE UPPER BOUNDARY FUNCTIONAL

Abstract

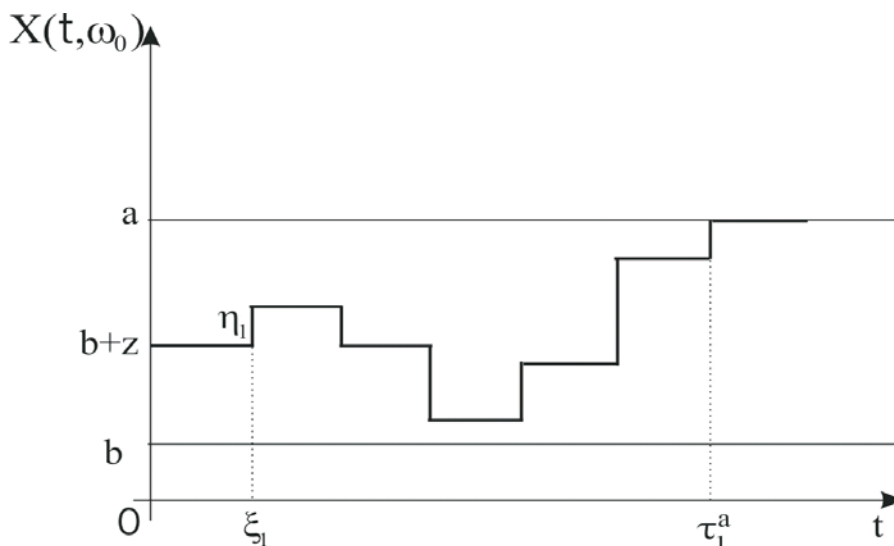
On the given two dimensional sequence of independent identically distributed random variables the semi-markov random walk with delaying screens in the "b" and "a" (b > 0) are constructed . Then the Laplace transformation of the first moment reaching the delaying screen in the "a" and its expectation and variance is obtained.

1.Problem statement. Let on a probability space $(\Omega, F, P(\cdot))$ be given a sequence of independent equally distributed random variables $\{\xi_i(\omega), \eta_i(\omega)\}$, $i = \overline{1, \infty}$ where $\xi_i(\omega) > 0$, $E\eta_1(\omega) > 0$ and the numbers a, b , and z , be given so that $b > 0$, $a > b$ and $0 < z \leq a - b$.

By these random variables we construct the process

$$X_1(t, \omega) = b + z + \sum_{i=1}^k \eta_i(\omega), \text{ if } \sum_{i=1}^k \xi_i(\omega) \leq t < \sum_{i=1}^{k+1} \xi_i(\omega), \quad k = \overline{0, \infty}, \quad \sum_1^0 = 0$$

and call it a step-wise process of semi-markov walk.



The process constructed in such a way is called a semi-markov walk process with delaying screens "b" and "a"

$$X(t, \omega) = \zeta_k(\omega), \text{ if } \sum_{i=1}^k \xi_i(\omega) \leq t < \sum_{i=1}^{k+1} \xi_i(\omega), \quad k \geq 0, \quad \left(\sum_1^0 = 0 \right)$$

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$$\zeta_k(\omega) = \min\{a, \max\{b, \zeta_{k-1}(\omega) + \eta_k(\omega)\}\}, \quad k \geq 1,$$

$$\zeta_0(\omega) = b + z.$$

One of realizations of the semi-markov process $X(t, \omega)$ looks like

Assume

$$\tau_1^a(\omega) = \inf\{t : X(t, \omega) = a\}$$

and $\tau_1^a(\omega) = \infty$, if for all $t > 0$ $X(t, \omega) < a$.

We call the random variable $\tau_1^a(\omega)$ the first moment reaching the screen "a".

The goal of the paper is to find the Laplace transformation of distribution of the random variable $\tau_1^a(\omega)$ and its number characteristics .

A number of papers [1, 4] were devoted to probability distributions and number characteristics study of boundary functionals of the semi-markov walk. Explicit relation for distributions of the first moment reaching of the zero level by a random walk generated by symmetric continuously distributed variables was obtained in [1, p.362] . When the walk occurs by any distribution, analytic form of distribution of the first moment reaching of the zero level that is not convenient for application, was obtained in [4] .

2. Formulation and proof of the main results. Introduce the following denotation:

$$L(\theta/b + z) = E\left(e^{-\theta\tau_1^a} / X(\theta) = b + z\right) =$$

$$= \int_{z=0}^{\infty} e^{-\theta t} dP\{\tau_1^a(\omega) < t / X(0, \omega) = b + z\},$$

$$L(\theta) = Ee^{-\theta\tau_1^a(\omega)} = \int_{z=0}^{\infty} L(\theta, b + z) P\{X(0, \omega) \in dz\}, \quad \theta > 0,$$

$$\varphi(\theta) = Ee^{-\theta\xi_1(\omega)} = \int_{i=0}^{\infty} e^{-\theta t} dP\{\xi_1(\omega) < t\}, \quad \theta > 0.$$

Theorem. For the process $X(t, \omega)$ it holds

$$L(\theta/b + z) = \varphi(\theta) P\{\eta_1(\omega) > a - b - z\} + \varphi(\theta) L(\theta/b) P\{\eta_1(\omega) < -z\} +$$

$$+ \varphi(\theta) \int_{y=b}^a L(\theta/y) d_y P\{\eta_1(\omega) < y - b - z\}$$

Proof. First we notice that by the theorem from [5] the first moment reaching of $\tau_1^a(\omega)$ is the eigen random value i.e. $P(\tau_1^a(\omega) < \infty) = 1$.

By total probability relation, we have

$$P\{\tau_1^a(\omega) > t / X(0, \omega) = b + z\} = P\{\xi_1(\omega) > t\} +$$

$$+ \int_{s=0}^t \int_{y=b}^a P\{\tau_1^a(\omega) > t - s / X(0, \omega) = y\} \times$$

$$\times P\{\xi_1(\omega) \in ds; b+z+\eta_1(\omega) < a; \max[b, b+z+\eta_1(\omega)] \in dy\}.$$

By means of this equality it is easy to get the following integral equation with respect to the conditional Laplace transformation $L(\theta/b+z)$:

$$L(\theta/b+z) = \varphi(\theta) P\{\eta_1(\omega) > a-b-z\} + \varphi(\theta) L(\theta/b) P\{\eta_1(\omega) < -z\} + \\ + \varphi(\theta) \int_{y=b}^a L(\theta/y) dy P\{\eta_1(\omega) < y-b-z\} \quad (1)$$

We can solve equation (1) for arbitrarily distributed random variable $\eta_1(\omega)$ by the sequential approximations method, but such a solution is not suitable for applications. We'll show that equation (1) admits exact solution.

Corollary1. Let $\eta_1(\omega)$ have Laplace distribution with parameters λ and μ . Then the Laplace conditional transformation will be of the form:

$$L(\theta/b+z) = \frac{\mu\varphi(\theta) [\lambda - k_2(\theta)] k_2(\theta) e^{k_1(\theta)(b+z)}}{[\lambda - k_2(\theta)]^2 k_2(\theta) e^{k_1(\theta)a} - [\lambda - k_1(\theta)]^2 k_1(\theta) e^{-[k_2(\theta)-k_1(\theta)]b+k_2(\theta)a}} - \\ - \frac{\mu\varphi(\theta) [\lambda - k_1(\theta)] k_1(\theta) e^{k_2(\theta)(b+z)}}{[\lambda - k_2(\theta)]^2 k_2(\theta) e^{[k_2(\theta)-k_1(\theta)]b+k_1(\theta)a} - [\lambda - k_1(\theta)]^2 k_1(\theta) e^{k_2(\theta)a}}.$$

Proof. Let η_1 have the Laplace distribution

$$P\{\eta_1 < x\} = \begin{cases} \frac{\lambda}{\lambda + \mu} e^{\mu x}, & x < 0, \mu > 0, \\ 1 - \frac{\mu}{\lambda + \mu} e^{-\lambda x}, & x > 0, \lambda > 0, \end{cases}$$

$$P_\eta(x) = \begin{cases} \frac{\lambda\mu}{\lambda + \mu} e^{\mu x}, & x < 0, \mu > 0, \\ \frac{\lambda\mu}{\lambda + \mu} e^{-\lambda x}, & x > 0, \lambda > 0. \end{cases}$$

Then equation (1) will be of the form:

$$L(\theta/b+z) = \frac{\mu\varphi(\theta)}{\lambda + \mu} e^{-\lambda(a-b-z)} + \frac{\lambda\varphi(\theta)}{\lambda + \mu} e^{-\mu z} L(\theta/b) + \\ + \frac{\lambda\mu\varphi(\theta)}{\lambda + \mu} e^{\lambda(b+z)} \int_{y=b+z}^a L(\theta/y) e^{-\lambda y} dy + \frac{\lambda\mu\varphi(\theta)}{\lambda + \mu} e^{-\mu(b+z)} \int_{y=b}^{b+z} L(\theta/y) e^{\mu y} dy \quad (2)$$

From this integral equation we get a differential equation with constant coefficients:

$$L_a''(\theta/b+z) - (\lambda - \mu) L_a'(\theta/b+z) - \lambda\mu [1 - \varphi(\theta)] L_a(\theta/b+z) = 0$$

that has the solution

$$L_a(\theta/b+z) = c_1(\theta) e^{k_1(\theta)(b+z)} + c_2(\theta) e^{k_2(\theta)(b+z)}, \quad (3)$$

$k_i(\theta)$, $i = 1, 2$ are the roots of the characteristic equation

$$k^2(\theta) - (\lambda - \mu)k(\theta) - \lambda\mu[1 - \varphi(\theta)] = 0.$$

Now, we are to find $c_i(\theta)$, $i = 1, 2$. To this end, from integral equation (2) we find the following initial conditions:

$$\left. \begin{aligned} L(\theta/a) &= \frac{\mu\varphi(\theta)}{\lambda + \mu} + \frac{\lambda\varphi(\theta)}{\lambda + \mu} e^{-\mu(a-b)} L(\theta/b) + \\ &+ \frac{\lambda\mu\varphi(\theta)}{\lambda + \mu} e^{-\mu a} \int_{y=b}^a L(\theta/y) e^{\mu y} dy, \\ L'(\theta/a) &= \frac{\lambda\mu\varphi(\theta)}{\lambda + \mu} + \frac{\lambda\mu\varphi(\theta)}{\lambda + \mu} e^{-\mu(a-b)} L(\theta/b) - \\ &- \frac{\lambda\mu^2\varphi(\theta)}{\lambda + \mu} e^{-\mu a} \int_{y=b}^a L(\theta/y) e^{\mu y} dy. \end{aligned} \right\}$$

On the other hand, from the solution of differential equation (3) we find

$$\begin{aligned} L_a(\theta/a) &= c_1(\theta) e^{k_1(\theta)a} + c_2(\theta) e^{k_2(\theta)a}, \\ L'_a(\theta/a) &= c_1(\theta) k_1(\theta) e^{k_1(\theta)a} + c_2(\theta) k_2(\theta) e^{k_2(\theta)a}. \end{aligned}$$

Hence we find the following system of algebraic inhomogeneous equations for $c_i(\theta)$, $i = 1, 2$.

$$\left. \begin{aligned} &\{\mu[\lambda - k_2(\theta)] e^{k_1(\theta)a} - [\lambda - k_1(\theta)] k_1(\theta) e^{-\mu(a-b)+k_1(\theta)b}\} c_1(\theta) + \\ &+ \{\mu[\lambda - k_1(\theta)] e^{k_2(\theta)a} - [\lambda - k_2(\theta)] k_2(\theta) e^{-\mu(a-b)+k_2(\theta)b}\} c_2(\theta) = \mu^2\varphi(\theta), \\ &\{\lambda[\lambda - k_2(\theta)] e^{k_1(\theta)a} + [\lambda - k_1(\theta)] k_1(\theta) e^{-\mu(a-b)+k_1(\theta)b}\} c_1(\theta) + \\ &+ \{\lambda[\lambda - k_1(\theta)] e^{k_2(\theta)a} + [\lambda - k_2(\theta)] k_2(\theta) e^{-\mu(a-b)+k_2(\theta)b}\} c_2(\theta) = \lambda\mu\varphi(\theta). \end{aligned} \right\}$$

From this system, by Kramer's relation, we find

$$\begin{aligned} c_1(\theta) &= \frac{\mu\varphi(\theta) [\lambda - k_2(\theta)] k_2(\theta)}{[\lambda - k_2(\theta)]^2 k_2(\theta) e^{k_1(\theta)a} - [\lambda - k_1(\theta)]^2 k_1(\theta) e^{-[k_2(\theta)-k_1(\theta)]b+k_2(\theta)a}}, \\ c_2(\theta) &= -\frac{\mu\varphi(\theta) [\lambda - k_1(\theta)] k_1(\theta)}{[\lambda - k_2(\theta)]^2 k_2(\theta) e^{[k_2(\theta)-k_1(\theta)]b+k_1(\theta)a} - [\lambda - k_1(\theta)]^2 k_1(\theta) e^{k_2(\theta)a}}. \end{aligned}$$

We substitute the found $c_1(\theta)$, $c_2(\theta)$ in (3) and get

$$\begin{aligned} L(\theta/b+z) &= \frac{\mu\varphi(\theta) [\lambda - k_2(\theta)] k_2(\theta) e^{k_1(\theta)(b+z)}}{[\lambda - k_2(\theta)]^2 k_2(\theta) e^{k_1(\theta)a} - [\lambda - k_1(\theta)]^2 k_1(\theta) e^{-[k_2(\theta)-k_1(\theta)]b+k_2(\theta)a}} - \\ &- \frac{\mu\varphi(\theta) [\lambda - k_1(\theta)] k_1(\theta) e^{k_2(\theta)(b+z)}}{[\lambda - k_2(\theta)]^2 k_2(\theta) e^{[k_2(\theta)-k_1(\theta)]b+k_1(\theta)a} - [\lambda - k_1(\theta)]^2 k_1(\theta) e^{k_2(\theta)a}}. \end{aligned}$$

Corollary 2. Let $\eta_1(\omega)$ have the Laplace distribution with parameters λ and μ , moreover $\lambda < \mu$. Then

$$E\tau_1^a(\omega) = \left\{ \frac{\lambda^3}{\mu(\lambda - \mu)^2} + \frac{\lambda\mu}{\lambda - \mu} (a - b) - \frac{\lambda^2}{\mu(\lambda - \mu)} e^{-\mu(a-b)} + \right.$$

$$\begin{aligned}
 & + \left. \frac{\mu}{\lambda - \mu} e^{-\lambda(a-b)} - \frac{\lambda^3}{\mu(\lambda - \mu)^2} e^{-(\lambda-\mu)(a-b)} \right\} E\xi_1(\omega), \\
 D\tau_1^a(\omega) = & \left\{ \frac{\lambda^3}{\mu(\lambda - \mu)^2} + \frac{\lambda\mu}{\lambda - \mu} (a - b) - \frac{\lambda^2}{\mu(\lambda - \mu)} e^{-\mu(a-b)} + \right. \\
 & \left. + \frac{\mu}{\lambda - \mu} e^{-\lambda(a-b)} - \frac{\lambda^3}{\mu(\lambda - \mu)^2} e^{-(\lambda-\mu)(a-b)} \right\} D\xi_1(\omega) + \\
 & + \left\{ -\lambda^3 \frac{\lambda^3 + 3\lambda^2\mu - 2\lambda\mu^2 - 5\mu^3}{(\lambda - \mu)^4} + \frac{\lambda\mu(\lambda^2 + \mu^2)}{(\lambda - \mu)^2} (a - b) + \right. \\
 & + \frac{2\lambda^2}{(\lambda - \mu)^3} e^{-(\lambda+\mu)(a-b)} + \frac{\lambda^6}{\mu^2(\lambda - \mu)^4} e^{2(\lambda-\mu)(a-b)} - \frac{\lambda^4}{\mu^2(\lambda - \mu)^2} e^{2\mu(a-b)} - \\
 & - \frac{\mu^2}{(\lambda - \mu)^2} e^{-2\lambda(a-b)} + \left[\frac{\mu(\lambda^2 + \mu^2) - 2\lambda(\lambda^2 + 2\mu^2)}{(\lambda - \mu)^3} - \frac{2\lambda\mu^2}{(\lambda - \mu)^2} (a - b) \right] e^{-\lambda(a-b)} + \\
 & + \left[\frac{2\lambda^2(\lambda^3 + \mu^3) - \lambda^3\mu(5\lambda + 4\mu)}{\mu^2(\lambda - \mu)^3} - \frac{2\lambda^3}{(\lambda - \mu)^2} (a - b) \right] e^{-\mu(a-b)} + \\
 & \left. + \left[\frac{2\lambda^3(\lambda + \mu)^2 - (\lambda^2 + \mu^2)}{\mu(\lambda - \mu)^4} - \frac{4\lambda^4}{(\lambda - \mu)^3} (a - b) \right] e^{-(\lambda-\mu)(a-b)} \right\} [E\xi_1(\omega)]^2
 \end{aligned}$$

Proof. Obviously

$$\begin{aligned}
 L_a(\theta) &= \int_{z=0}^{a-b} L_a(\theta/b + z) dP\{X\{0, \omega\} < b + z\} = \\
 &= \int_{z=0}^{a-b} L_a(\theta/b + z) dP\{\min(a, b + \eta_1^+(\omega)) < b + z\} = L_a(\theta/a) P\{\eta_1^+(\omega) > a - b\} - \\
 & - \int_{z=0}^{a-b} L_a(\theta/b + z) dP\{\eta_1^+(\omega) > z\} = - \sum_{i=1}^2 \frac{k_i(\theta)}{\lambda - k_i(\theta)} c_i(\theta) e^{\lambda b - [\lambda - k_1(\theta)]a} + \\
 & + \sum_{i=1}^2 \frac{\lambda}{\lambda - k_i(\theta)} c_i(\theta) e^{k_i(\theta)b}. \tag{4}
 \end{aligned}$$

Now, let's find $E\tau_1^a(\omega)$, $D\tau_1^a(\omega)$.

By the Laplace transformation property we have

$$E\tau_1^a(\omega) = -L'(0) \text{ and } D\tau_1^a(\omega) = L''(0) - [L'(0)]^2$$

Then from (4) we get

$$E\tau_1^a(\omega) = \left\{ \frac{\lambda^3}{\mu(\lambda - \mu)^2} + \frac{\lambda\mu}{\lambda - \mu} (a - b) - \frac{\lambda^2}{\mu(\lambda - \mu)} e^{-\mu(a-b)} + \right.$$

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$$\begin{aligned}
& + \frac{\mu}{\lambda - \mu} e^{-\lambda(a-b)} - \frac{\lambda^3}{\mu(\lambda - \mu)^2} e^{-(\lambda - \mu)(a-b)} \Big\} E\xi_1(\omega), \\
D\tau_1^a(\omega) = & \left\{ \frac{\lambda^3}{\mu(\lambda - \mu)^2} + \frac{\lambda\mu}{\lambda - \mu} (a - b) - \frac{\lambda^2}{\mu(\lambda - \mu)} e^{-\mu(a-b)} + \right. \\
& \left. + \frac{\mu}{\lambda - \mu} e^{-\lambda(a-b)} - \frac{\lambda^3}{\mu(\lambda - \mu)^2} e^{-(\lambda - \mu)(a-b)} \right\} D\xi_1(\omega) + \\
& + \left\{ -\lambda^3 \frac{\lambda^3 + 3\lambda^2\mu - 2\lambda\mu^2 - 5\mu^3}{(\lambda - \mu)^4} + \frac{\lambda\mu(\lambda^2 + \mu^2)}{(\lambda - \mu)^2} (a - b) + \right. \\
& + \frac{2\lambda^2}{(\lambda - \mu)^3} e^{-(\lambda + \mu)(a-b)} + \frac{\lambda^6}{\mu^2(\lambda - \mu)^4} e^{2(\lambda - \mu)(a-b)} - \frac{\lambda^4}{\mu^2(\lambda - \mu)^2} e^{2\mu(a-b)} - \\
& - \frac{\mu^2}{(\lambda - \mu)^2} e^{-2\lambda(a-b)} + \left[\frac{\mu(\lambda^2 + \mu^2) - 2\lambda(\lambda^2 + 2\mu^2)}{(\lambda - \mu)^3} - \frac{2\lambda\mu^2}{(\lambda - \mu)^2} (a - b) \right] e^{-\lambda(a-b)} + \\
& + \left[\frac{2\lambda^2(\lambda^3 + \mu^3) - \lambda^3\mu(5\lambda + 4\mu)}{\mu^2(\lambda - \mu)^3} - \frac{2\lambda^3}{(\lambda - \mu)^2} (a - b) \right] e^{-\mu(a-b)} + \\
& \left. + \left[\frac{2\lambda^3(\lambda + \mu)^2 - (\lambda^2 + \mu^2)}{\mu(\lambda - \mu)^4} - \frac{4\lambda^4}{(\lambda - \mu)^3} (a - b) \right] e^{-(\lambda - \mu)(a-b)} \right\} [E\xi_1(\omega)]^2.
\end{aligned}$$

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