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MODELING OF BELLMAN-HARRIS BRANCHING PROCESSES

Abstract

In the present paper the Bellman-Harris branching random processes are investigated. Different properties of these processes are investigated and on the basis of obtained above theoretical results the noted processes are modeled in computer.

Recently many papers appear which is devoted to stochastic analysis of branching random processes. In this aspect in [1] the mathematical model of Galton-Watson classical random process is constructed and this process is modeled in computer.

Further on the base of modeling the basic probability characteristics of indicated process are investigated.

In the present paper the analogous investigations are carried out for Bellman-Harris branching random processes. The mathematical construction of probability space on which the Bellman-Harris branching random processes are determined, is stated in detail in [2].

The Bellman-Harris processes differ from the Galton-Watson process only by the fact that the moment of appearance of particle up to its conversion or disappearance passes some random time, named as life time. Rather, assume, that the process begins from one particle of zero age having random time of life l with distribution of the probability $G(t) = P\{l \leq t\}$. At the end of life the particle with the probabilities p_k , k = 0, 1, 2, ... is converted to k new particles of zero age. It is assumed, that the distribution of life time is the same for all particles and the probabilities p_k doesn't depend on the age of particles at the moment of its conversion and from the number of other existent particles. The process continues until there exists even one particle.

Thus the Bellman-Harris branching processes can be determined by the distribution function of life time G(t) and the generating function

$$h\left(z\right) = \sum_{k=0}^{x} p_k z^k$$

of the number of direct descendants of one particle.

The Bellman-Harris process whose life time is identically equal to unit will be the Galton-Watson process. Therefore any Bellman-Harris branching random process is closely concerned with the corresponding Galton-Watson process, whose generating function of number of direct descendant of one particle is h(z).

Let $\xi(t)$, $t \ge 0$ be Bellman-Harris branching process with generating function of number of particles at the moment of time t

$$F(t;z) = \sum_{k=0}^{\infty} P\left\{\xi(t) = k\right\} z^{k} = M\left\{z^{\xi(t)} | \xi(0) = 1\right\}$$

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It is evident [2], that the generating function F(t; z) satisfies the following nonlinear integral equation

$$F(t;z) = z \left[1 - G(t)\right] + \int_{0}^{t} h\left(F(t - u; z)\right) dG(u).$$
(1)

If we search all possible functions for distribution life time of one particle G(t)and all possible generating functions h(z) (the generating functions of corresponding Galton-Watson processes), then among the obtained Bellman-Harris branching processes may be found such ones whose number of particles in every fixed moment of time t > 0 with positive probability is equal to infinity. If there is no such phenomenon, then the process is called regular. The continuity in zero of distribution function of the life time G(t), i.e., G(0+) = 0 and the finiteness $m = \Sigma k p_k$ of average number of direct descendants of one particle are sufficient conditions of regularity.

As it is noted any Bellman-Harris branching process is closely connected with the corresponding Galton-Watson process. The degeneration of Bellman-Harris process is not connected generally with the life time of particle and consequently, the probability q of degeneration of the Belman-Kharis process is equal to the nearest o solution z = h(z) in unit interval $0 \le z \le l$.

The generating function F(t; z) is connected with double inequalities with the generating function $h_n(z)$ of the number of particles in the *n*-th generation or the corresponding Galton-Watson process. Namely, if $0 \le z \le q$, then

$$h_{n}(z) - (q-z) \sum_{k=0}^{n-1} \left[h'_{n}(q) \right]^{k} \left\{ G^{k^{*}}(t) - G^{k+1^{*}}(t) \right\} \leq \\ \leq F(t;z) \leq h_{n}(z) + (q-z) \left[h'_{n}(q) \right]^{n} G^{n^{*}}(t)$$

Here the natural number n and nonnegative number t are arbitrary, $G^{k^*}(t)$ - kmultiple convolution of the distribution function G with itself. In the critical case (m = h(1) = 1) this inequalities take the especially simple form. If m = 1, then it
is evident, that q = 1 consequently,

$$h_n(z) - (1-z) \left[1 - G^{n^*}(t) \right] \le F(t;z) \le h_n(z) + (1-z) G^{n^*}(t)$$
(2)

for all $0 \le z \le 1$, $n = 0, 1, 2, \dots$ $t \ge 0$.

In [3] using the double inequality (2) the limit theorem on convergence of normed Bellman-Harris processes to the Jirina critical branching process is proved (the branching process with continuous state space).

And in [4] applying the integral equation (1) the limit theorem on convergence of one-dimensional distributions of Bellman-Harrin branching processes to the Jirini one-dimensional distributions is proved.

In particular cases we can obtain integral equation of form (1) for the generating function F(t; z) by means of which the probabilistic characteristics of the process are investigated.

Note, that by obtaining integral equation (1) it is assumed, that the any particle at the expiry of its life time with the definite probability (p_k) is converted to the random number of new particles of zero age.

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If we assume, that the particle on ending of life time generates two new particles, then for the generating function F(t; z) the following integral equation is obtained

$$F(t;z) = z \left[1 - G(t)\right] + \int_{0}^{t} \left[F(t - u;z)\right]^{2} dG(u).$$
(3)

The evident generalization of this model is the assumption about that the particle at the expiry of its life time generates equally r new particles of the same type, where r is a fixed integer $r \ge 2$. In this case the integral equation (3) is substituted by the following one:

$$F(t;z) = z \left[1 - G(t)\right] + \int_{0}^{t} \left[F(t - u;z)\right]^{r} dG(u).$$
(4)

In general case the integral equations (1), (3) and (4) aren't solved. Therefore for the investigation of the described processes we model their behaviour in computer and mark the compare with the results obtained by an analytical way.

Exactly as in [1] we assume, that the process begins from one particle of the zero age $\xi(0) = 1$ and the probability law of development of population has the following form

$$p_k = P\left\{\xi_1 = k\right\} = \frac{1}{2^k}, \ k = 1, 2, \dots$$
 $\sum_{k=1}^{\infty} p_k = 1$

Besides, we assume, that the distribution function of the life time G(t) has exponential form, i.e.,

$$G(t) = 1 - e^{-\lambda t}, \quad t \ge 0$$

By composing a program for the modelling of behaviour of Bellman-Harris branching processes in computer it is chosen pseudorandom number α from the interval (0,1), and the parameter λ is chosen equal to 2. If α is situated in the interval $\left(\frac{1}{2^{k+1}}, \frac{1}{2^k}\right)$, i.e., $p_{k+1} < \alpha \le p_k$, then in the next generation the quantity of anew generated particles is equal to k. If $\frac{1}{2^k} < \alpha < 1$, then the number of particles in the next generation is equal to 0, i.e. the process is degenerated.

At this approach the following characteristics of Bellman-Harris branching processes was investigated:

-summary number of particles in the definite moment

- average number of particles in definite moment

- variance of amount of particles in definite moment.

It should be noted, that the indicated investigation is conducted for non-normalized processes. At this the 16 realizations of process are considered, and it is introduced only 6-th realization of the process in graphic for summary number of particles (fig.1).

Further, the analogous investigation is conducted for Bellman-Harris normalized processes. In this case the above mentioned characteristics of normalized processes are studied. The 16 realizations of the process are considered, but in graphics for summary number of particles (fig.4) it is introduced the 9-th realization of process. All results obtained at modelling in computer are shown in graphics and in tables. The summary numbers of particles in the system for non-normalized case (6-th realization) is shown in fig.1. The average and variance of amount of particles in

the system for non-normalized case is shown in fig.2 and 3. The corresponding characteristics for normalized processes is shown in fig. 5 and 6.

At modelling it was chosen the time of modelling T = 750. Unlike the case for Galton-Watson branching processes, here during the time T appears n = 380generations. It is explained by the randomness of life time of particle (but the generation takes some time unit).

According to the graphics of process behaviour and in tables obtained at modelling in computer we can argue, that the results obtained by the analogous way in [3] and [4] are reliable

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Fig. 1.

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Fig. 2.

Fig. 3.

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Fig. 5.

Fig. 4.

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Fig. 6.

The numb. of	Time	The results	mean	Dispersion
generation		of graphics	value	
		in fig.1		
1	0.118806	6	4.549	6.277
20	9.74541	36	55.726	66.295
40	22.0758	1	65.245	123.214
60	30.1153	22	46.415	71.488
80	39.7085	45	64.707	98.939
100	52.4288	4	100.861	134.168
120	62.9442	8	112.041	137.315
140	71.65	14	114.734	111.732
160	81.3515	20	102.106	97.489
180	90.9352	92	132.782	130.555
200	99.5314	213	12.203	122.492
220	109.667	22	73.505	79.951
240	118.557	11	43.961	45.932
260	128.221	1	43.387	56.352
280	133.987	22	58.961	58.282
300	145.564	4	66.400	79.9656
320	155.261	24	38.281	52.968
340	158.226	105	66.250	90.389
360	169.456	72	33.703	57.775
380	182.017	106	32.426	71.696

Tabl.1 (Non-normalized case)

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The numb. of	Time	The results	mean	Dispersion
generation		of graphics	value	
		in fig.4		
1	0.118806	1	1.290	1.901
20	9.74541	0.05	0.560	0.747
40	22.0758	0.225	0.984	1.143
60	27.6953	0.0166667	0.31	0.299
80	39.7085	0.1	0.373	0.422
100	52.4288	0.12	0.431	0.607
120	62.9442	0.183333	0.352	0.537
140	71.65	0.378571	0.317	0.336
160	81.3515	0.01875	0.154	0.231
180	90.9352	0.0611111	0.260	0.371
200	99.5314	0.25	0.210	0.327
220	109.667	0.213636	0.160	0.381
240	118.557	0.0291667	0.404	0.807
260	128.221	0.0461538	0.421	0.707
280	133.987	0.0142857	0.472	0.798
300	145.564	0.0666667	0.717	1.023
320	155.261	0.11875	0.622	0.944
340	162.845	0.0117647	0.587	0.786
360	170.187	0.0583333	0.481	0.884
380	180.479	0.142105	0.167	0.612

Tabl.2 (Normalized case)

References

[1]. Aliyev S.A., Gastratova M.G. *Modelling of Galton-Watson branching random processes.* Proceedings of IMM of NAS of Azerbaijan, 2003, v.XVIII (XXVI), pp.3-9. (Russian)

[2]. Harris T. Theory of branching random processes. M.: "Mir", 1966, 355p. (Russian)

[3]. Aliyev S.A. On one limit theorem for Bellman-Harris processes. In. book: Nekotoriye voprosi teorii sluchaynix processov, Kiev, 1984, pp.3-10. (Russian)

[4]. Aliyev S.A. On convergence of Bellman-Harris branching processes. Ukr. Math. Journal, v.42, No6, pp.827-832. (Russian)

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